



KENYATTA UNIVERSITY

UNIVERSITY EXAMINATIONS 2011/2012

SECOND SEMESTER EXAMINATION FOR THE DEGREE OF BACHELOR OF SCIENCE (STATISTICS AND COMPUTING)

SST 302: PROGRAMMING LANGUAGE FOR STATISTICS II

DATE: Tuesday 3rd April 2012

TIME: 8.00a.m – 10.00a.m

INSTRUCTIONS: Answer Question ONE [COMPULSORY] and any other TWO questions

1. (a) Define the following terms

i. Random variable

(1mk)

ii. Monte Carlo methods

(1mk)

iii. Bootstrap methods

(1mk)

(b) Given a Cauchy random variate X with a cumulative distribution function

$$F(x) = \left\{ \frac{1}{2} + \frac{1}{\pi} \arctan \left(\left(\frac{x - \mu}{\sigma} \right) \right) \right\} \quad 0 \leq \mu < \infty, \sigma^2 > 0$$

Give an R code for generating random variables using the inverse transform method

(5mks)

(c) Consider a geometric probability mass function given by

$$P(X = j) = \begin{cases} pq^{j-1} & \text{for } j = 1, 2, \dots \\ 0 & \text{Otherwise} \end{cases}$$

i. Obtain the cumulative distribution function for Geometric probability mass function above

(3mks)

ii. Give an algorithm for simulating random variates from the distribution function obtained in (i) above

(4mks)

iii. Give an R code to implement the algorithm in (ii) above

(3mks)

- (d) Using the recursive property of the binomial mass function, illustrate an algorithm for generating random variates from a binomial mass function given by

$$P(X = j) = \begin{cases} \frac{n!}{(n-j)!j!} p^j (1-p)^{n-j} & \text{for } j = 0, 1, 2, \dots, n \\ 0 & \text{Otherwise} \end{cases}$$

. Also show the *R* code to implement the algorithm illustrated.

(10mks)

- (e) Highlight any two advantages of bootstrap as a re-sampling technique to statistical inference

(2mks)

2. (a) Describe how to sample by rejection sampling from a density f using samples from a density g .

(3mks)

- (b) On average, what proportion of observations from g will be accepted as observations from density f ? When will rejection sampling be efficient.

(2mks)

- (c) Describe two advantages of rejection sampling over the inversion method for simulating from a distribution.

(2mks)

- (d) Let X have a probability density function of the form

$$f(x) \propto \begin{cases} 1 - \cos(x) & 0 \leq x \leq \frac{\pi}{2} \\ 0 & \text{Otherwise} \end{cases}$$

- i. Describe a rejection algorithm for simulating from the random variable X , based upon simulating over the rectangle $\left[0, \frac{\pi}{2}\right] \times [0, 1]$. Calculate the acceptance probability of the algorithm

(6mk)

- ii. Show how the method of inversion can be used to simulate from a random variable Y with

$$f(y) \propto \begin{cases} \left(\frac{y^2}{2}\right) & 0 \leq y \leq \frac{\pi}{2} \\ 0 & \text{Otherwise} \end{cases}$$

Note that for all $0 \leq x \leq \frac{\pi}{2}$, $1 - \cos(x) \leq \frac{x^2}{2}$. Use this fact to describe a rejection algorithm to simulate from X using Y . Calculate the acceptance probability of the rejection algorithm.

(7mks)

3. (a) For X , $E[X] = (8 - 4\pi + \pi^2)/(4(\pi - 2))$ and $E[X^2] = (48 - 6\pi^2 + \pi^3)/(12(\pi - 2))$. Given that $\text{cov}(\exp(X), X), \text{cov}(\exp(X), X^2) > 0$, outline how control variates, using $E[X]$ and $E[X^2]$, can be used to estimate $\theta = E[\exp(X)]$, based upon a sample $x_1, x_2, \dots, x_n$ from X

(6mks)

- (b) i. Describe the bootstrap as a way of estimating the statistical properties of a statistic $\varphi(X_1, X_2, \dots, X_n)$ for an independent random sample $X_1, X_2, \dots, X_n$. In particular, give expressions for the bootstrap estimate of the variance and bias

(8mks)

- ii. For a random variable X , let $\theta = E[X]$. Suppose that $x_1, x_2, \dots, x_n$ is a random (independent) sample from X . Then $\hat{\theta} = \bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$ is unbiased estimate of $E[X]$. Show that the variance of the bootstrap estimate of θ given the random sample $x_1, x_2, \dots, x_n$ is $\frac{1}{n^2} \sum_{i=1}^n (x_i - \bar{x})^2$.

(6mks)

4. (a) A time series x is assumed to come from a first order auto-regressive model

$$X_i = \rho X_{i-1} + \epsilon_i$$

where $2 \leq i \leq n$ for an unknown constant ρ . However the distribution of the errors $\{\epsilon_i\}$ is completely unknown except that it is reasonable to assume they are independent and identically distributed with mean 0

- i. Find an expression for $\hat{\rho}$ where $\hat{\rho}$ is the least squares estimate of ρ

(5mks)

- ii. Explain how you would simulate a bootstrap sample from the distribution of $\hat{\rho}$

(5mks)

- iii. How would you use this bootstrap sample to test the hypothesis that $\rho = 0.5$? Give an outline of the procedure to be used.

(5mks)

- (b) Give explicit details of how to simulate by inversion of the cumulative distribution function (cdf) from a random variable Z with probability density function,

$$g(z) = \begin{cases} \left(\frac{1215}{z^6} \right) & \text{for } (z > 3) \\ 0 & \text{otherwise.} \end{cases}$$

(5mks)

5. (a) Suppose that $Y_1, \dots, Y_n$ is a random sample from a population and that θ is a population parameter that we would like to estimate. We have two potential estimators, $\hat{\theta}_1 = u_1(Y_1, \dots, Y_n)$ and $\hat{\theta}_2 = u_2(Y_1, \dots, Y_n)$. Carefully describe a simulation experiment to compare the performance of $\hat{\theta}_1$ and $\hat{\theta}_2$ as estimators of θ . You may assume that, for a given θ , we can generate (pseudo-) random instances from the population distribution. Are there any graphical techniques that you think might be useful in comparing pieces of output from the simulation?

(10mks)

(b) Consider a simple regression model

$$Y_t = \beta x_t + \varepsilon_t, \varepsilon_t \sim N(0, \sigma^2) \quad (1)$$

where $t = 1, \dots, n$ and the ε_t are independent.

- i. Write an R function that takes three arguments, the slope parameter β , the error standard deviation σ , and a vector of values of the explanatory variable x . The default values for β and σ should both be 1. The function should return simulated values $y_1, \dots, y_n$ from model (1) where n is the length of the explanatory variable vector. (2mks)
- ii. Write an R function that uses the function in part (i) above and returns a vector of r instances of the parameter estimate $\hat{\beta}$. (3mks)
- iii. Write an R function whose inputs include a vector of possible β values and a vector of possible σ values and, for every possible combination simulates instances of $\hat{\beta}$ and outputs the simulated approximation to the mean square error. It should be clear from the output which pair of β and σ values each MSE value corresponds to (5mks)